# Permutation Excess Entropy and Mutual Information between the Past and Future

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Abstract We address the excess entropy, which is a measure of complexity for stationary time series, from the ordinal point of view. We show that the permutation excess entropy is equal to the mutual information between two adjacent semi-infinite blocks in the space of orderings for finite-state stationary ergodic Markov processes. This result may spread a new light on the relationship between complexity and anticipation.

**Keywords** : Permutation Entropy, Excess Entropy, Mutual Information, Complexity, Duality

## 1 Introduction

Recently, it was found that much of the information contained in stationary time series can be captured by orderings of values, not the values themselves [1]. The permutation entropy rate which was first introduced in [5, 6] quantifies the average uncertainty of orderings of values per time unit. This is in contrast to the usual entropy rate which quantifies the average uncertainty of values per time unit. However, surprisingly, it is known that the permutation entropy rate is equal to the entropy rate for finite-state stationary stochastic processes [1, 2]. Similar results for dynamical systems are also known [2, 3, 6, 15, 18].

In our previous work [14], we found a new proof of the equality between the permutation entropy rate and the entropy rate based on the duality between values and orderings, which can be seen as a Galois connection [11] (categorical adjunction [17] for partially ordered sets, however, we do not refer to the Galois connection explicitly in this paper). By making use of the duality, we also proved that the permutation excess entropy is equal to the excess entropy for finite-state stationary ergodic Markov processes.

The excess entropy has attracted interest from the complex systems community for decades [4, 7, 9, 10, 12, 13, 16, 19]. By definition, the excess entropy is the sum of entropy over-estimates over finite length of words [10]. However, it can be expressed as the mutual information between the past and future, namely, the

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International Journal of Computing Anticipatory Systems, Volume 26, 2014 Edited by D. M. Dubois, CHAOS, Liège, Belgium, ISSN 1373-5411 ISBN 2-930396-15-6 mutual information between two adjacent semi-infinite blocks of stochastic variables. Thus, the excess entropy can be interpreted as a measure of global correlation present in a system.

In this paper, based on the duality between values and orderings, we show that the permutation excess entropy also admit the mutual information expression in the space of orderings when the process is finite-state stationary ergodic Markov one. This result partially justifies the claim that the permutation excess entropy measures global correlation at the level of orderings of values present in stationary time series.

This paper is organized as follows. In Section 2, we review the duality between values and orderings. In Section 3, we explain the permutation excess entropy. In Section 4, we present a proof of the claim that the permutation excess entropy has the mutual information expression for finite-state stationary ergodic Markov processes. In Section 5, we give conclusions.

#### 2 Duality between Values and Orderings Explained

Let  $A_n = \{1, 2, \dots, n\}$  be a finite alphabet consisting of natural numbers from 1 to n. We consider  $A_n$  as a totally ordered set ordered by the usual 'less-than-or-equal-to' relationship.

We denote the set of all permutations of length  $L \ge 1$  by  $S_L$ . Namely, each element  $\pi \in S_L$  is a bijection on the set of indexes  $\{1, 2, \dots, L\}$ . For convenience, we denote each permutation  $\pi \in S_L$  by the string  $\pi(1) \cdots \pi(L)$ .

we denote each permutation  $\pi \in S_L$  by the string  $\pi(1) \cdots \pi(L)$ . For each word  $s_1^L := s_1 \cdots s_L := (s_1, \cdots, s_L) \in A_n^L = \underbrace{A_n \times \cdots \times A_n}_L$  of length

 $L \geq 1$ , we define its permutation type  $\pi \in S_L$  by re-ordering symbols  $s_1, \dots, s_L$  in increasing order:  $s_1^L$  is of type  $\pi$  if we have  $s_{\pi(i)} \leq s_{\pi(i+1)}$  and  $\pi(i) < \pi(i+1)$  when  $s_{\pi(i)} = s_{\pi(i+1)}$  for  $i = 1, 2, \dots, L-1$ . For example,  $\pi(1)\pi(2)\pi(3)\pi(4) = 3142$  for  $s_1^4 = 2312$  because  $s_3s_1s_4s_2 = 1223$ .

We introduce the map  $\phi : A_n^L \to S_L$  that sends each word  $s_1^L$  to its unique permutation type  $\pi = \phi(s_1^L)$ . This map  $\phi$  classifies or coarse-grains words of length L by the criterion whether they have the same permutation type. In general,  $\phi$ is many-to-one map. For example, all of 111, 112, 122, 222  $\in A_2^3$  have the same permutation type  $\pi \in S_3$  defined by  $\pi(1)\pi(2)\pi(3) = 123$  (identity on  $\{1, 2, 3\}$ ).

Now, we list the properties of the map  $\phi$  which will be used later.

**Lemma 1** For  $s_1^L, t_1^L \in A_n^L$ ,  $\phi(s_1^L) = \phi(t_1^L)$  if and only if  $s_k \leq s_j \Leftrightarrow t_k \leq t_j$  for all  $1 \leq j \leq k \leq L$ .

*Proof.* See Corollary 4 in [14].

**Lemma 2** Let  $n \ge i \ge 1$ . Fix  $\pi \in S_L$ . Assume that there is no  $s_1^L \in A_{i-1}^L$  such that  $\phi(s_1^L) = \pi$ , but there exists  $s_1^L \in A_i^L$  such that  $\phi(s_1^L) = \pi$  (When i = 1 we define  $A_{i-1} = A_0 = \emptyset$ ). Then,

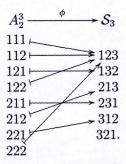
- (i) there exists a unique  $s_1^L \in A_i^L$  such that  $\phi(s_1^L) = \pi$ . Moreover, if  $\phi(t_1^L) = \pi$  for  $t_1^L \in A_n^L$ , then there exist  $c_1, \dots, c_L$  such that  $s_k + c_k = t_k$  for  $k = 1, \dots, L$ and  $0 \le c_{\pi(1)} \le \dots \le c_{\pi(L)} \le n - i$ .
- (ii)  $|\phi^{-1}(\pi)| = {\binom{L+n-i}{n-i}}$ , where |X| denotes the cardinality of a set X.

*Proof.* See Lemma 5 in [14]. (ii) follows from the fact that the number of sequences  $a_1 \cdots a_L$  satisfying  $0 \le a_1 \le a_2 \le \cdots \le a_L \le n-i$  is given by the binomial coefficient  $\binom{L+n-i}{n-i}$ .

For example, let  $\phi \in S_5$  be given by  $\pi(1)\pi(2)\pi(3)\pi(4)\pi(5) = 24315$ . We have  $\phi(s_1^5) = \pi$  for  $s_1^5 = s_1s_2s_3s_4s_5 = 31213 \in A_3^5$ . Consider  $t_1^5 = t_1t_2t_3t_4t_5 = 41325 \in A_5^5$  and  $c_1c_2c_3c_4c_5 = 10112$ . We have  $\phi(t_1^5) = \pi$  and  $t_2t_4t_3t_1t_5 = 12345 = 11233 + 01112 = s_2s_4s_3s_1s_5 + c_2c_4c_3c_1c_5$ .

As a more thorough illustration of Lemma 2, let us write down how  $\phi$  sends each word to its permutation type for L = 3 and n = 1, 2.

When n = 1, the unique element  $111 \in A_1^3$  is mapped to  $123 \in S_3$ . When n = 2, we have



For example, there is no  $s_1^3 \in A_1^3$  such that  $\phi(s_1^3) = 132 \in S_3$ . On the other hand,  $\phi^{-1}(132) = \{121\}$  for  $\phi: A_2^3 \to S_3$ . We have  $\phi^{-1}(123) = \{111, 112, 122, 222\}$  for  $\phi: A_2^3 \to S_3$ . Note that  $|\phi^{-1}(123)| = 4 = \binom{3+2-1}{2-1}$ .

Let us introduce the map  $\mu : S_L \to \mathbb{N}^L$ , where  $\mathbb{N} = \{1, 2, \dots\}$  is the set of all natural numbers, by the following procedure:

(i) given a permutation π ∈ S<sub>L</sub>, we decompose the sequence π(1) ··· π(L) into maximal ascending subsequences. A subsequence i<sub>j</sub> ··· i<sub>j+k</sub> of a sequence i<sub>1</sub> ··· i<sub>L</sub> is called a maximal ascending subsequence if it is ascending, namely, i<sub>j</sub> ≤ i<sub>j+1</sub> ≤ ··· ≤ i<sub>j+k</sub>, and neither i<sub>j-1</sub>i<sub>j</sub> ··· i<sub>j+k</sub> nor i<sub>j</sub>i<sub>j+1</sub> ··· i<sub>j+k+1</sub> is ascending.

(ii) If  $\pi(1) \cdots \pi(i_1)$ ,  $\pi(i_1+1) \cdots \pi(i_2)$ ,  $\cdots$ ,  $\pi(i_{k-1}+1) \cdots \pi(L)$  is the decomposition of  $\pi(1) \cdots \pi(L)$  into maximal ascending subsequences, then we define the word  $s_1^L \in \mathbb{N}^L$  by

$$s_{\pi(1)} = \dots = s_{\pi(i_1)} = 1, s_{\pi(i_1+1)} = \dots = s_{\pi(i_2)} = 2, \dots,$$
$$s_{\pi(i_{k-1})+1} = \dots = s_{\pi(L)} = k.$$

(iii) We define  $\mu(\pi) = s_1^L$ .

By construction, we have  $\phi \circ \mu(\pi) = \pi$  when  $\mu(\pi) \in A_n^L$  for all  $\pi \in \mathcal{S}_L$ .

For example, the decomposition of  $15423 \in S_5$  into maximal ascending subsequences is 15, 4, 23. We obtain  $\mu(\pi) = s_1 s_2 s_3 s_4 s_5 = 13321$  by putting  $s_1 s_5 s_4 s_2 s_3 = 11233$ .

The map  $\mu$  can be seen as the dual to the map  $\phi$  in the following sense:

Theorem 3 Let us put

$$X = \{s_1^L \in A_n^L | \phi^{-1}(\pi) = \{s_1^L\} \text{ for some } \pi \in \mathcal{S}_L\},$$
(1)

$$Y = \{\pi \in \mathcal{S}_L ||\phi^{-1}(\pi)| = 1\}.$$
 (2)

Then,  $\phi$  restricted on X is a map into Y,  $\mu$  restricted on Y is a map into X, and they form a pair of mutually inverse maps. Furthermore, we have

$$X = \{s_1^L \in A_n^L | \text{for any } 1 \le i \le n-1 \text{ there exist } 1 \le j < k \le L \\ \text{such that } s_j = i+1, s_k = i\}$$
(3)

*Proof.* See Theorem 9 in [14].

For the map  $\phi: A_2^3 \to \mathcal{S}_3$ , the duality

$$X \xrightarrow{\phi}_{\longleftarrow} Y$$

is given by

121	~~ 132
211	~~ 213
212	~~ 231
221	→ 312.

(4)

#### **3** Permutation Excess Entropy

Let  $\mathbf{S} = \{S_1, S_2, \dots\}$  be a finite-state stationary stochastic process, where each stochastic variable  $S_i$  takes its value in  $A_n$ . By stationarity, we mean

$$\Pr\{S_1 = s_1, \cdots, S_L = s_L\} = \Pr\{S_{k+1} = s_1, \cdots, S_{k+L} = s_L\}$$

for any  $k, L \ge 1$  and  $s_1, \dots, s_L \in A_n$ . Hence, we can define the probability of the occurrence of each word  $s_1^L \in A_n^L$  by  $p(s_1^L) := p(s_1 \dots s_L) := \Pr\{S_1 = s_1, \dots, S_L = s_L\}$ .

The entropy rate  $h(\mathbf{S})$  of a finite-state stationary stochastic process  $\mathbf{S}$ , which quantifies the average uncertainty of values per time unit, is defined by

$$h(\mathbf{S}) = \lim_{L \to \infty} \frac{1}{L} H(S_1^L), \tag{5}$$

where  $H(S_1^L) = H(S_1, \dots, S_L) = -\sum_{s_1^L \in A_n^L} p(s_1^L) \log_2 p(s_1^L)$ . The limit exists for any finite-state stationary stochastic process [8].

The *permutation entropy rate* quantifies the average uncertainty of orderings of values per time unit. It is defined by

$$h^*(\mathbf{S}) = \lim_{L \to \infty} \frac{1}{L} H^*(S_1^L) \tag{6}$$

if the limit exists, where  $H^*(S_1^L) = H^*(S_1, \dots, S_L) = -\sum_{\pi \in S_L} p(\pi) \log_2 p(\pi)$  and  $p(\pi)$  is the probability that  $\pi$  is realized in **S**, namely,  $p(\pi) = \sum_{s_1^L \in \phi^{-1}(\pi)} p(s_1^L)$  for  $\pi \in S_L$ .

**Theorem 4** For any finite-state stationary stochastic process S, the permutation entropy rate  $h^*(S)$  exists and

$$h^*(\mathbf{S}) = h(\mathbf{S}). \tag{7}$$

*Proof.* The proof appealing to ergodic theory is found in [1, 2]. For an alternative proof based on the duality between values and orderings, see [14].

The entropy rate can be seen as a measure of randomness of a finite-state stationary stochastic process. Meanwhile the excess entropy can be interpreted as a measure of complexity [12]. More precisely, it measures global correlation present in a system. The *excess entropy*  $\mathbf{E}(\mathbf{S})$  of a finite-state stationary stochastic process  $\mathbf{S}$  is defined by [10]

$$\mathbf{E}(\mathbf{S}) = \lim_{L \to \infty} \left( H(S_1^L) - h(\mathbf{S})L \right)$$
(8)

if the limit exists. If  $\mathbf{E}(\mathbf{S})$  exists, then we have [10]

$$\mathbf{E}(\mathbf{S}) = \sum_{L=1}^{\infty} \left( H(S_L | S_1^{L-1}) - h(\mathbf{S}) \right) = \lim_{L \to \infty} I(S_1^L; S_{L+1}^{2L}),$$
(9)

where H(Y|X) is the conditional entropy of Y given X and I(X;Y) is the mutual information between X and Y for stochastic variables X and Y.

The permutation excess entropy was introduced in [14] by imitating the definition of the excess entropy. The permutation excess entropy  $\mathbf{E}^*(\mathbf{S})$  of a finite-state stationary stochastic process  $\mathbf{S}$  is defined by

$$\mathbf{E}^{*}(\mathbf{S}) = \lim_{L \to \infty} \left( H^{*}(S_{1}^{L}) - h^{*}(\mathbf{S})L \right), \tag{10}$$

if the limit exists. However, it is unclear what form of correlation the permutation excess entropy quantifies from this expression. In the following discussion, we partially resolve this problem. We will show that the equality

$$\mathbf{E}^*(\mathbf{S}) = \lim_{L \to \infty} I(\phi(S_1^L); \phi(S_{L+1}^{2L}))$$
(11)

holds for any finite-state stationary ergodic Markov process S. Recall that the entropy rate and the excess entropy of a finite-state stationary Markov process S are given by

$$h(\mathbf{S}) = -\sum_{i,j=1}^{n} p_i p_{ij} \log_2 p_{ij}$$

and

$$\mathbf{E}(\mathbf{S}) = -\sum_{i=1}^{n} p_i \log_2 p_i + \sum_{i,j=1}^{n} p_i p_{ij} \log_2 p_{ij},$$

respectively, where  $P = (p_{ij})$  is the transition matrix and  $\mathbf{p} = (p_1, \dots, p_n)$  is the stationary distribution. P and  $\mathbf{p}$  satisfy the following properties:

- (i)  $p_{ij} \ge 0$  for all  $1 \le i, j \le n$ .
- (ii)  $\sum_{i=1}^{n} p_{ij} = 1$  for all  $1 \le i \le n$ .
- (iii)  $p_i \ge 0$  for all  $1 \le i \le n$ .

(iv) 
$$\sum_{i=1}^{n} p_i = 1.$$

(v)  $\sum_{i=1}^{n} p_i p_{ij} = p_j$  for all  $1 \le j \le n$ .

 $p_{ij}$  is the probability of transition from state *i* to state *j*.  $p_i$  is the probability of observing state *i*. Thus, the probability of the occurrence of each word  $s_1^L \in A_n^L$  is given by  $p(s_1^L) = p_{s_1} p_{s_1 s_2} \cdots p_{s_{L-1} s_L}$ . A finite-state stationary Markov process **S** is *ergodic* if and only if its transition matrix *P* is *irreducible* [20]: a matrix *P* is *irreducible* if for all  $1 \leq i, j \leq n$  there exists l > 0 such that  $p_{ij}^{(l)} > 0$ , where  $p_{ij}^{(l)}$  is the (i, j)-th element of  $P^l$ . For an irreducible non-negative matrix *P*, stationary distribution  $\mathbf{p} = (p_1, \cdots, p_n)$  exists uniquely and satisfies  $p_i > 0$  for all  $1 \leq i \leq n$ .

In our previous work [14], we showed that the equality

 $\mathbf{E}^*(\mathbf{S}) = \mathbf{E}(\mathbf{S}) \tag{12}$ 

holds for any finite-state stationary ergodic Markov process. The key point of the proof is that the probability

$$q_{L} = \sum_{\substack{\pi \in \mathcal{S}_{L}, \\ |\phi^{-1}(\pi)| > 1}} p(\pi) = \sum_{\pi \notin Y} p(\pi)$$
(13)

diminishes exponentially fast as  $L \to \infty$  for any finite-state stationary ergodic Markov process, where the set Y is given by (2) in Theorem 3. For the proof of the equality (11), we also appeal to this fact. Hence, we shortly review the reason why this fact follows.

Let L be a positive integer. We introduce the following probability  $\beta_s$  for each symbol  $s \in A_n$ :

$$\beta_s = \Pr\{s_1^N | s_j \neq s \text{ for any } 1 \le j \le N\},\tag{14}$$

where  $N = \lfloor L/2 \rfloor$  and  $\lfloor x \rfloor$  is the largest integer not greater than x.

**Lemma 5 (Lemma 12 in [14])** Let **S** be a finite-state stationary stochastic process and  $\epsilon$  be a positive real number. If  $\beta_s \leq \epsilon$  for any  $s \in A_n$ , then  $q_L \leq 2n\epsilon$ .

*Proof.* We shall prove  $\sum_{\pi \in Y} p(\pi) \ge 1 - 2n\epsilon$ , where the set Y is given by (2) in Theorem 3. Let us consider words  $s_1^L \in A_n^L$  satisfying the following two conditions:

(i) Each symbol  $s \in A_n$  appears in  $s_1^N$  at least once.

(ii) Each symbol  $s \in A_n$  appears in  $s_{N+1}^L$  at least once.

By the assumption of the lemma, we have

$$\Pr\{s_1^N | (i) \text{ holds}\} \ge 1 - n\epsilon,$$

because

$$\Pr\{s_1^N | (i) \text{ holds}\} + \sum_{s=1}^n \Pr\{s_1^N | s_j \neq s \text{ for any } 1 \le j \le N\} \ge 1.$$

Similarly,

 $\Pr\{s_{N+1}^L|(ii) \text{ holds}\} \ge 1 - n\epsilon$ 

holds because of the stationarity. Hence, we obtain

 $\Pr\{s_1^L | \text{both (i) and (ii) hold}\} \ge 1 - 2n\epsilon.$ 

Since any word  $s_1^L \in A_n^L$  satisfying both (i) and (ii) is a member of the set X given by (1) in Theorem 3, we obtain

$$\sum_{\pi \in Y} p(\pi) = \sum_{s_1^L \in X} p(s_1^L) \ge \Pr\{s_1^L | \text{both (i) and (ii) hold}\} \ge 1 - 2n\epsilon.$$

Let **S** be a finite-state stationary ergodic Markov process whose transition matrix is P and stationary distribution is **p**. We can write  $\beta_s$  in the following form by using Markov property:

$$\beta_s = \sum_{\substack{s_j \neq s, \\ 1 \le j \le N}} p(s_1 \cdots p_N) = \sum_{\substack{s_j \neq s, \\ 1 \le j \le N}} p_{s_1} p_{s_1 s_2} \cdots p_{s_{N-1} s_N} = \langle (P_s)^{N-1} \mathbf{u}_s, \mathbf{p} \rangle, \quad (15)$$

where the matrix  $P_s$  is defined by

$$(P_s)_{ij} = \begin{cases} 0 & \text{if } i = s \\ p_{ij} & \text{otherwise,} \end{cases}$$

the vector  $\mathbf{u}_s = (u_1, \dots, u_n)$  is defined by  $u_i = 0$  if i = s otherwise  $u_i = 1$  and  $\langle \dots, \dots \rangle$  is the usual inner product in the *n*-dimensional Euclidean space.

We can prove that the non-negative largest eigenvalue  $\lambda$  of  $P_s$  is strictly less than 1 and absolute value of any other eigenvalue of  $P_s$  is not greater than  $\lambda$  by using Perron-Frobenius Theorem for non-negative matrices and the irreducibility of P (Lemma 13 in [14]). Hence, by decomposing  $P_s$  into a sum of a diagonalizable matrix and a nilpotent matrix, we obtain the following lemma:

**Lemma 6** Let **S** be a finite-state stationary ergodic Markov process. There exists  $0 \leq \alpha < 1, C > 0$  and a positive integer k such that  $\beta_s \leq C\alpha^L L^k$  for any  $s \in A_n$  and sufficiently large L.

# 4 Mutual Information Expression of Permutation Excess Entropy

In this section, we give a proof of the equality (11) for finite-state stationary ergodic Markov processes. We make use of the notions of *rank sequences* and *rank variables* which are introduced in [2].

Rank sequences of length L are words  $r_1^L \in \mathbb{N}^L$  satisfying  $1 \leq r_i \leq i$  for  $i = 1, \dots, L$ . We denote the set of all rank sequences of length L by  $\mathcal{R}_L$ . Clearly,  $|\mathcal{R}_L| = L! = |\mathcal{S}_L|$ .

We can transform each word  $s_1^L \in A_n^L$  into a rank sequence  $r_1^L \in \mathcal{R}_L$  by defining

$$r_i = \sum_{j=1}^{i} \delta(s_j \le s_i), \ i = 1, \cdots, L,$$
 (16)

where  $\delta(X) = 1$  if the proposition X is true, otherwise  $\delta(X) = 0$ . Namely,  $r_i$  is the number of indexes j  $(1 \le j \le i)$  such that  $s_j \le s_i$ . Thus, we obtain the map  $\varphi: A_n^L \to \mathcal{R}_L$  such that  $\varphi(s_1^L) = r_1^L$ .

We can show that the map  $\varphi : A_n^L \to \mathcal{R}_L$  is compatible with the map  $\phi : A_n^L \to \mathcal{S}_L$ . Namely, there exists a bijection  $\iota : \mathcal{R}_L \to \mathcal{S}_L$  satisfying  $\iota \circ \varphi = \phi$  [14].

Given a stationary stochastic process  $\mathbf{S} = \{S_1, S_2, \dots\}$ , its associated rank variables are defined by  $R_i = \sum_{j=1}^n \delta(S_j \leq S_i)$  for  $i = 1, 2, \dots$ . Note that rank variables  $R_i$   $(i = 1, 2, \dots)$  are not stationary stochastic variables in general. By the compatibility between  $\phi$  and  $\varphi$ , we have

$$H(R_1^L) = H^*(S_1^L) = H(\phi(S_1^L))$$
(17)

for  $L \geq 1$ .

Now, let **S** be a finite-state stationary ergodic Markov process. By (12), we know that the permutation excess entropy  $\mathbf{E}^*(\mathbf{S})$  exists. By (17) and the chain rule, we have

$$\mathbf{E}^{*}(\mathbf{S}) = \lim_{L \to \infty} \left( H^{*}(S_{1}^{L}) - h^{*}(\mathbf{S})L \right) = \lim_{L \to \infty} \left( H(R_{1}^{L}) - h^{*}(\mathbf{S})L \right)$$
(18)  
$$= \sum_{L=1}^{\infty} \left( H(R_{L}|R_{1}^{L-1}) - h^{*}(\mathbf{S}) \right).$$

Since the infinite sum in (19) converges, we obtain

$$\left| H(R_{L+1}^{2L}|R_1^L) - h^*(\mathbf{S})L \right| = \left| \sum_{i=1}^{L} \left( H(R_{L+i}|R_1^{L+i-1}) - h^*(\mathbf{S}) \right) \right| \xrightarrow[L \to \infty]{} 0.$$
(19)

By the definition of mutual information, we have

$$I(\phi(S_1^L);\phi(S_{L+1}^{2L})) = H(\phi(S_{L+1}^{2L})) - H(\phi(S_{L+1}^{2L})|\phi(S_1^L)).$$

In addition, we have  $H(\phi(S_{L+1}^{2L})) = H(\phi(S_1^L)) = H^*(S_1^L)$  by the stationarity of **S**. Hence, it is sufficient to show that

$$\left| H(\phi(S_{L+1}^{2L}) | \phi(S_1^L)) - h^*(\mathbf{S}) L \right| \underset{L \to \infty}{\to} 0$$
(20)

to prove the equality (11). However, by (19), this reduces to showing that

$$\left| H(\phi(S_{L+1}^{2L}) | \phi(S_1^L)) - H(R_{L+1}^{2L} | R_1^L) \right| \underset{L \to \infty}{\to} 0,$$
(21)

which is equivalent to showing that

$$|H(\phi(S_1^L), \phi(S_{L+1}^{2L})) - H(\phi(S_1^{2L}))| \xrightarrow[L \to \infty]{} 0$$
 (22)

by (17).

**Lemma 7** For  $s_1^{2L}, t_1^{2L} \in A_n^{2L}$ , if  $\phi(s_1^{2L}) = \phi(t_1^{2L})$ , then  $\phi(s_1^L) = \phi(t_1^L)$  and  $\phi(s_{L+1}^{2L}) = \phi(t_{L+1}^{2L})$ . Namely, the partition of  $A_n^{2L}$  by the map  $\phi: A_n^{2L} \to S_{2L}$  is a refinement of the partition of  $A_n^L \times A_n^L = A_n^{2L}$  by the map  $\phi \times \phi: A_n^L \times A_n^L \to S_L \times S_L$ .

Proof. The claim follows immediately from Lemma 1.

Lemma 8

$$0 \leq H(\phi(S_1^{2L})) - H(\phi(S_1^L), \phi(S_{L+1}^{2L}))$$

$$\leq \left( \sum_{\substack{\pi', \pi'' \in S_L, \\ |\phi^{-1}(\pi')| > 1 \text{ or } |\phi^{-1}(\pi'')| > 1}} p(\pi', \pi'') \right) 2n \log_2(L+n)$$
(23)

holds for any finite-state stationary stochastic process S, where

$$p(\pi', \pi'') = \sum_{\substack{s_1^L \in \phi^{-1}(\pi'), \\ s_{L+1}^{2L} \in \phi^{-1}(\pi'')}} p(s_1^{2L})$$

for  $\pi', \pi'' \in S_L$ .

Proof. By Lemma 7, we can write

$$\begin{split} &H(\phi(S_1^{2L})) - H(\phi(S_1^L), \phi(S_{L+1}^{2L})) \\ &= -\sum_{\pi \in \mathcal{S}_{2L}} p(\pi) \log_2 p(\pi) + \sum_{\pi', \pi'' \in \mathcal{S}_L} p(\pi', \pi'') \log_2 p(\pi', \pi'') \\ &= \sum_{\pi', \pi'' \in \mathcal{S}_L} \left( -\sum_{\substack{\phi^{-1}(\pi) \subseteq \\ (\phi \times \phi)^{-1}(\pi', \pi'')}} p(\pi) \log_2 p(\pi) + p(\pi', \pi'') \log_2 p(\pi', \pi'') \right) \\ &= \sum_{\pi', \pi'' \in \mathcal{S}_L} \left( -\sum_{\substack{\phi^{-1}(\pi) \subseteq \\ (\phi \times \phi)^{-1}(\pi', \pi'')}} p(\pi) \log_2 p(\pi) + \sum_{\substack{\phi^{-1}(\pi) \subseteq \\ (\phi \times \phi)^{-1}(\pi', \pi'')}} p(\pi) \log_2 p(\pi', \pi'') \right) \\ &= \sum_{\substack{\pi', \pi'' \in \mathcal{S}_L, \\ p(\pi', \pi'') > 0}} p(\pi', \pi'') \left( -\sum_{\substack{\phi^{-1}(\pi) \subseteq \\ (\phi \times \phi)^{-1}(\pi', \pi'')}} \frac{p(\pi)}{p(\pi', \pi'')} \log_2 \frac{p(\pi)}{p(\pi', \pi'')} \right). \end{split}$$

By Lemma 2 (ii), we have

$$0 \le -\sum_{\substack{\phi^{-1}(\pi) \subseteq \\ (\phi \times \phi)^{-1}(\pi',\pi'')}} \frac{p(\pi)}{p(\pi',\pi'')} \log_2 \frac{p(\pi)}{p(\pi',\pi'')} \le 2n \log_2(L+n).$$

If  $|\phi^{-1}(\pi')| = 1$  and  $|\phi^{-1}(\pi'')| = 1$  hold for  $(\pi', \pi'') \in S_L \times S_L$ , then  $|(\phi \times \phi)^{-1}(\pi', \pi'')| = 1$ . In this case, if  $p(\pi', \pi'') > 0$ , then we have

$$-\sum_{\substack{\phi^{-1}(\pi)\subseteq\\ (\phi\times\phi)^{-1}(\pi',\pi'')}}\frac{p(\pi)}{p(\pi',\pi'')}\log_2\frac{p(\pi)}{p(\pi',\pi'')}=0.$$

Lemma 9 (22) holds for any finite-state stationary ergodic Markov process S.

Proof. We have

$$\sum_{\substack{\pi',\pi''\in\mathcal{S}_L,\\ |\phi^{-1}(\pi')|>1 \text{ or } |\phi^{-1}(\pi'')|>1}} p(\pi',\pi'') \leq \sum_{\substack{|\phi^{-1}(\pi')|>1,\\ \pi''\in\mathcal{S}_L}} p(\pi',\pi'') + \sum_{\substack{|\phi^{-1}(\pi'')|>1,\\ \pi''\in\mathcal{S}_L}} p(\pi',\pi'') = 2q_L.$$

By Lemma 5 and Lemma 6, there exist  $0 \le \alpha < 1$ , C > 0 and k > 0 such that  $q_L \le C\alpha^L L^k$  for sufficiently large L if **S** is a finite-state stationary ergodic Markov process. The claim follows from Lemma 8.

Thus, we get our main theorem in this paper:

**Theorem 10** The equality (11)

$$\mathbf{E}^*(\mathbf{S}) = \lim_{L \to \infty} I(\phi(S_1^L); \phi(S_{L+1}^{2L}))$$

holds for any finite-state stationary ergodic Markov process S.

### 5 Conclusions

In this paper, we showed that the permutation excess entropy is equal to the mutual information between the past and future in the space of orderings for finite-state stationary ergodic Markov processes.

Combining this result with the equality between the excess entropy and the permutation excess entropy for finite-state stationary ergodic Markov processes proved in our previous work [14], we can see that the global correlation measured by the mutual information between the past and future in a given finite-state stationary ergodic Markov process is fully captured by a different coding of the process, namely, coding by the orderings of values.

We hope that our result gives rise to a new insight into the relationship between complexity and anticipation.

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